

A Topological Approach of Principal Component Analysis

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Abstract: Large datasets are increasingly widespread in many disciplines. The exponential growth of data requires the development of more data analysis methods in order to process information more efficiently. In order to better visualize the data, many methods such as Principal Component Analysis (PCA) and MultiDimensional Scaling (MDS) allow to extract a low-dimensional structure from high-dimensional data set. The proposed approach, called Topological Principal Component Analysis (TPCA), is a multidimensional descriptive method which studies a homogeneous set of continuous variables defined on the same set of individuals. It is a topological method of data analysis that consists of comparing and classifying proximity measures from among some of the most widely used proximity measures for continuous data. Proximity measures play an important role in many areas of data analysis, the results strongly depend on the proximity measure chosen. So, among the many existing measures, which one is most useful? Are they all equivalent? How to identify the one that is most appropriate to analyze the correlation structure of a set of quantitative variables. TPCA proposes an appropriate adjacency matrix associated to an unknown proximity measure according to the data under consideration, then analyzes and visualizes, with graphic representations, the relationship structure of the variables relating to, the well known PCA problem. It uses the concept of neighborhood graphs and compares a set of proximity measures for continuous data which can be more-or-less equivalent a topological equivalence criterion between two proximity measures is defined and statistically tested according to the topological correlation between the variables considered. An example on real data illustrates the proposed approach.

Keywords: Proximity Measure, Neighborhood Graph, Adjacency Matrix, Topological Equivalence, Correlation Matrix, MDS Graphical Representation

1. Introduction

Choosing a proximity measure from among the many available measures greatly influences the results of any data analysis method, moreover, these measures are more-or-less equivalent according to the concept of the neighborhood graph structure used.

A topological equivalence criterion is defined between proximity measures from the topological structure induced by each measure.

Large datasets are increasingly common and are often difficult to interpret. Principal component analysis (PCA) [16, 10, 5, 18] is a technique for reducing the dimensionality of such datasets, increasing interpretability but at the same time minimizing information loss. It is an exploratory tool for continuous data.

PCA is an adaptive technique for continuous data,

variants of this technique have been developed and tailored to various different data types and structures.

In order to suitably interpret the large datasets, methods are needed are required to drastically reduce their dimensionality in an interpretable way. Many techniques have been developed for this purpose, but PCA is one of the oldest and most widely used. PCA is statistically considered as a widely used multivariate method for dimension reduction and as a technique of representing data. It aims to find common factors, the so-called principal components, in form of linear combinations of the variables under investigation. It allows to have an idea of the correlations structure of the set of variables, as well as possible similarities of behavior between individuals.

In the context of artificial intelligence, we often compare

des situations représentées par un ensemble d'objets, pour cela, nous devons choisir et spécifier la mesure de proximité entre les objets. Le contexte d'étude, le type de données et d'autres facteurs peuvent nous aider à choisir la mesure de proximité qui pourrait être appropriée. Cependant, le nombre de mesures possibles peut être assez grand.

De plus, ces mesures qui sont toujours possibles, sont-elles toutes équivalentes? Existe-t-il une mesure plus spécifique ou plus adaptée qu'une autre pour l'étude considérée? En matière de recherche d'information, le choix de la mesure de proximité est un problème essentiel sur lequel les résultats dépendent.

La présente étude propose un nouveau cadre pour comparer les mesures de proximité afin d'identifier celles qui sont similaires, ainsi, nous n'aurons plus besoin d'essayer toutes les mesures.

Ces comparaisons sont clarifiées par une mesure de proximité qui évalue la similarité ou la dissimilarité entre deux objets au sein d'un ensemble. Les mesures de proximité ont des propriétés mathématiques et des axiomes bien spécifiques.

La meilleure mesure est sélectionnée en fonction de la structure de l'ensemble de variables quantitatives à synthétiser, l'objectif est d'établir une PCA topologique. Les résultats de l'ATPCA sont différents selon la mesure de proximité sélectionnée.

Plusieurs auteurs ont étudié l'équivalence topologique des mesures de proximité, dans un cadre général [4, 17, 13, 24], dans le contexte de l'analyse discriminante [3] et de l'analyse de correspondance [2, 1], mais aucune dans le contexte de la PCA. Ainsi, dans cet article, nous montrons comment construire la matrice d'adjacence induite par une mesure de proximité inconnue mais qui prend en compte la structure topologique des variables que nous souhaitons décrire.

Dans cet article, nous comparons différentes mesures de proximité dans un but de synthétiser les relations d'un ensemble de variables continues dans le contexte topologique. Les comparaisons de ces mesures montrent que les résultats sont différents et dépendent de la mesure de proximité choisie. Le reste de l'article est organisé comme suit. Dans la section 2, nous discutons l'équivalence topologique entre deux mesures de proximité et montrons comment construire une matrice d'adjacence associée à une mesure de proximité, comment comparer et tester statistiquement le degré d'équivalence topologique entre les mesures de proximité et comment sélectionner la meilleure mesure pour décrire

topologiquement la structure des corrélations des variables. La section 3 présente un exemple illustratif et des mesures de proximité existantes sur des données continues et présente une comparaison entre elles. Cette comparaison aide les chercheurs à prendre rapidement une décision sur la mesure à utiliser pour les données considérées. Une conclusion de ce travail est donnée dans la section 4.

Le tableau 8 de l'Annexe résume certaines mesures de proximité classiques utilisées pour des données continues [23], nous donnons sur \mathbb{R}^n la définition de 15 d'entre elles.

Nous supposons que nous avons à notre disposition $\{x^k; k=1, \dots, p\}$ un ensemble de p variables quantitatives homogènes mesurées sur n individus. L'intérêt est d'analyser la structure topologique de toutes ces variables.

2. Topological Correlation

Le concept d'équivalence topologique entre deux mesures de proximité est basé sur le concept du graphe de voisinage. Deux mesures sont dites topologiquement équivalentes si leurs graphes induits sur l'ensemble d'objets restent identiques. Mesurer la similarité entre deux mesures de proximité revient à mesurer la similarité de leurs graphes de voisinage.

Considérons un ensemble $E = \{x^1, x^2, \dots, x^k, \dots, x^p\}$ de p objets dans \mathbb{R}^n , associés avec les p variables quantitatives.

Étant donnée une mesure de proximité notée u , nous pouvons définir une relation binaire de voisinage sur $E \times E$ notée V_u . Ainsi, nous pouvons construire un graphe de voisinage sur un ensemble d'objets-variables, où les sommets sont les variables et les arêtes sont définies par la propriété de la relation de voisinage. Il s'agit d'une matrice symétrique binaire.

Plusieurs définitions de graphes sont possibles pour construire cette matrice binaire. On peut choisir l'Arbre Spanning Minimal (MST) [11], le Graphe de Gabriel (GG) [15] ou, comme c'est le cas ici, le Graphe de Voisinage Relatif (RNG) [21].

Ainsi, étant donnée une mesure de proximité u , nous pouvons associer une matrice d'adjacence V_u d'ordre p , où toutes les paires (x^k, x^l) de variables voisines satisfont l'expression RNG suivante :

$$V_u(x^k, x^l) = \begin{cases} 1 & \text{if } u(x^k, x^l) \leq \max[u(x^k, x^r), u(x^r, x^l)] \\ & \forall x^k, x^l, x^r \in E, x^r \neq x^k \text{ and } x^r \neq x^l \\ 0 & \text{otherwise} \end{cases}$$

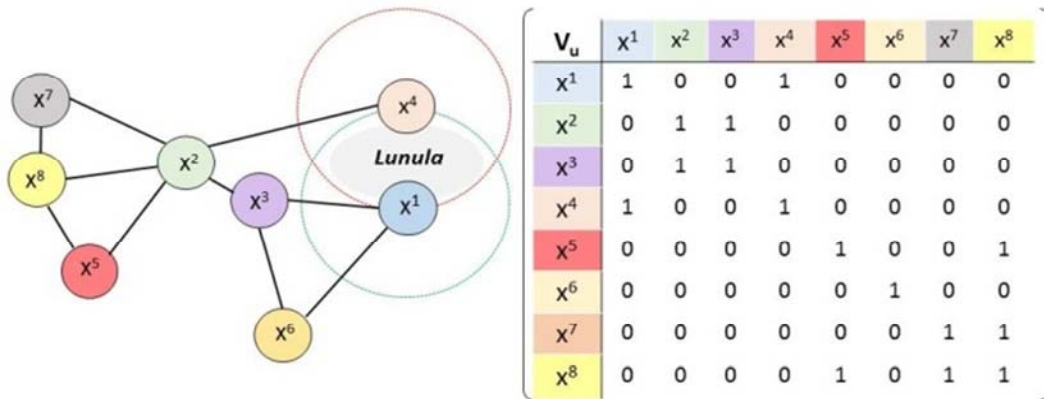


Figure 1. RNG example with eight variables - Adjacency matrix.

This means that if two variables x^k and x^l which verify the RNG property are connected by an edge, the vertices x^k and x^l are neighbors.

Thus, for any proximity measure given, u , we can associate an adjacency matrix V_u , of binary and symmetrical order p . Figure 1 illustrates an example of RNG in R^2 of a set of $p=8$ objects-variables.

For example, for the first and four variables, $V_u(x^1, x^4)=1$, it means that on the geometrical plane, the hyper-Lunula (intersection between the two hyperspheres centered on the two variables x^1 and x^4) is empty.

For a given neighborhood property (MST, GG or RNG), each measure u generates a topological structure on the objects in E which are totally described by the adjacency binary matrix V_u . In this paper, we chose to use the Relative Neighbors Graph (GNR).

2.1. Comparison and Selection of Proximity Measures

First we compare different proximity measures according to their topological similarity in order to regroup them and to better visualize their resemblances.

To measure the topological equivalence between two proximity measures u_i and u_j , we propose to test if the associated adjacency matrices V_{u_i} and V_{u_j} are different or not. The degree of topological equivalence between two proximity measures is measured by the following definition of concordance. The topological equivalence between two adjacency matrices satisfy the following expression:

$$S(V_{u_i}, V_{u_j}) = \frac{1}{p^2} \sum_{k=1}^p \sum_{l=1}^p \delta_{kl}(x^k, x^l)$$

$$\text{with } \delta_{kl}(x^k, x^l) = \begin{cases} 1 & \text{if } V_{u_i}(x^k, x^l) = V_{u_j}(x^k, x^l) \\ 0 & \text{otherwise} \end{cases}$$

Then, in our case, we want to compare these different proximity measures according to their topological equivalence in a context of correlation. So we define a criterion for measuring the deviation from the independence position.

The data can arise from several different sampling frameworks, and the interpretation of the hypothesis of no association depends on the framework. The question of interest is whether there is correlation between the two variables.

We construct the adjacency matrix denoted by V_{u^*} , which corresponds to the correlation matrix.

Thus, to examine the correlation structure between the variables, we examine the significance of their linear correlation coefficient. This adjacency matrix can be written as follows using the t-test of the linear correlation coefficient ρ of Bravais-Pearson. The adjacency matrix V_{u^*} associated to reference measure u^* satisfy the following expression:

$$V_{u^*}(x^k, x^l) = \begin{cases} 1 & \text{if } p\text{-value} = P[|T_{n-2}| > t\text{-value}] \leq \alpha \\ 0 & \text{otherwise} \end{cases} \quad \forall k=1, p, \forall l=1, p$$

Where p-value is the significance test of the correlation coefficient for the two-sided test of the null and alternative hypotheses, $H_0: \rho(x^k, x^l)=0$ vs. $H_1: \rho(x^k, x^l) \neq 0$.

The p-value is the evidence against a null hypothesis. The smaller the p-value, the stronger the evidence that you should reject the null hypothesis which means that there is no correlation between x^k and x^l variables in the population.

Formula for the Student t-test for significance of correlation: $t = \sqrt{(n-2)} (1 - r^2)$ with $v = n - 2$ degrees of freedom (d.f.) and $r = r(x^k, x^l)$ is the linear correlation coefficient observed between the variables x^k and x^l .

Let T_{n-2} be a t-distributed random variable of Student with $v = n - 2$ d.f. In this case, the null hypothesis is rejected with a p-value less or equal a chosen α significance level, for example $\alpha=5\%$. Using linear correlation test, if the p-value be very small, it means that there is very small opportunity that null hypothesis is correct, and consequently we can reject it. Statistical significance in statistics is achieved when a p-value is less than the significance level of α . The p-value is the probability of obtaining results which acknowledge that the null hypothesis is true.

The robustness according to the α error risk chosen for the null hypothesis, no linear correlation, can be studied by setting a minimum threshold in order to analyze the sensitivity of the results. Certainly the numerical results will change, but probably not their interpretation.

The binary and symmetric adjacency matrix build V_{u^*} , is associated with an unknown proximity measure denoted u^* and called a reference measure. Thus, with this reference proximity measure we can establish (V_{u_i}, V_{u^*}) , the topological equivalence between the two proximity measures u_i and u^* , by measuring the percentage of similarity between the adjacency matrix V_{u_i} and the reference adjacency matrix V_{u^*} .

In order to graphically describe the similarities between proximity measures, we can for example apply the notion of them scope [12], which is a methodological sequence of a clustering method on the results of a factorial method. In this case, a Principal Component Analysis (PCA) followed by a Hierarchical Ascendant Classification (HAC) were performed upon the 15 component dissimilarity matrix defined by: $[D]_{ij} = D(V_{u_i}, V_{u_j}) = 1 - S(V_{u_i}, V_{u_j})$ to partition them into homogeneous groups and to view their similarities.

We can use any classic visualization techniques to achieve this. For example, we can build a dendrogram of hierarchical clustering of the proximity measures. We can also use multidimensional scaling or any other technique, such as Laplacian projection, to map the 15 proximity measures into a two dimensional space.

Finally, in order to evaluate and determine the closest class of proximity measures to the reference measure u^* , we project the latter as a supplementary element into the two data analysis methods, positioned by the dissimilarity vector with 15 components $[D]_{*j} = 1 - S(V_{u^*}, V_{u_j})$.

2.2. Statistical Comparisons Between Proximity Measures

In this section, we use the Fisher's Exact Test [9] which is an alternative to the Chi-square test when the samples are small. The principle of this test is to determine if the configuration observed in the contingency table is an extreme situation compared to the possible situations taking into account the marginal distributions. Fisher's exact test is an exact statistical test used for the analysis of contingency tables. It is a test qualified as exact because the probabilities can be calculated exactly rather than relying on an approximation which becomes correct only asymptotically as for the chi-square test used in the contingency tables.

It is not based on a test statistic whose law is known when n is large enough but it calculates, as its name suggests, the exact p -value directly. To test statistically the topological equivalence between two proximity measures, this non parametric test compares these measures based on their associated adjacency matrices. Two proximity measures are statistically in topological equivalence if the null hypothesis H_0 of independence is rejected.

The comparison between indices of proximity measures has also been studied by Demsar [7] and Schneider & Borlund [19, 20] from a statistical perspective. The authors proposed an approach that compares similarity matrices obtained by each proximity measure, using Mantel's test [14], in a pairwise manner.

Fisher's exact test is the statistical test best suited to compare matched binary data, the Cohen's Kappa test [6] also but it is in general an asymptotic test. The Kendall or Spearman coefficient compares matched continuous data. It makes it possible in this context to measure the agreement or the concordance of the binary values of two adjacency matrices associated with two proximity measures. The Fisher's exact test between two adjacency matrices evaluates the topological equivalence between their proximity measures.

Let V_{ui} and V_{uj} be adjacency matrices associated with two proximity measures u_i and u_j . To compare the degree of topological equivalence between these two measures, we propose to test if the associated adjacency matrices are statistically different or not, using a non-parametric test of paired data. These binary and symmetric matrices of order p , are unfolded in two vector-matched components, consisting of $p(p+1)/2$ values, the p diagonal values and the $p(p-1)/2$ values above or below the diagonal.

The degree of topological equivalence between two proximity measures is evaluated from the Fisher's exact test, computed on the 2×2 contingency table formed by the two binary vectors of order $p(p+1)/2$.

We also test the topological equivalence between each proximity measure $u_{i=1,15}$ and the reference measure u_* by comparing the adjacency matrices V_{ui} and V_{u*} .

2.3. Graphical Representations - Variables & Individuals

In order to represent graphically the possible

topological links between the p quantitative variables, we use MultiDimensional Scaling (MDS) which makes it possible to find, for any distance matrix (similarity or dissimilarity) of size $p \times p$, a set of p points identified by their Euclidean coordinates whose distance matrix is equal to or very close to the given distance matrix.

We carry out the classical MDS [5], namely factorial analysis on similarity V_{u*} or dissimilarity $D_{u*}=U - V_{u*}$ table, where $U=1_p \cdot 1_p$ is the $p \times p$ matrix of 1s and 1_p denotes the p indicator vector of 1s.

The TPCA approach consist to perform the standardized PCA of the triple $\{V_{u*}; M; D_p\}$, where, V_{u*} is the adjacency matrix associated with the proximity measure u_* , the most appropriate measure for the considered data, $M=I_p$ is the identity matrix of order p and $D_p=1/p I_p$ is the weighted diagonal matrix of variable weights.

The TPCA can be performed from any adjacency matrix V_{ui} associated with each of the 15 proximity measures u_i considered. Aid for the interpretation of TPCA results are those of PCA. Graphical representations on factorial plans allow to visualize and identify the topological structure of the variables. As in PCA, for representations of variables, we consider the most significant variables on the axes, that is the variables highly correlated with factors, having a strong contribution and a good quality of representation, measured by the square cosine of the angle between main axes and initial axes.

For representations of active individuals, these are projected as illustrative elements. The quality of representation of these individuals on the factorial axes is measured by their squared cosine.

3. Illustrative Example and Empirical Results

To illustrate the TPCA, we use Eurostat data [8] on government finance of the 28 European Union (EU) countries in 2017. We examine how key government finance statistics have developed in the EU-28. Specifically, it considers general government gross debt, deficit/surplus, total revenue and total expenditure. Simple statistics of the considered variables are displayed in Table 1.

Table 1. Summary statistics of public finances.

Variable	N	Mean	Std. Dev.	Coef. Var. (%)	Min	Max
Debt	28	68.04	36.5	53.70	8.7	176.1
Deficit	28	-0.26	1.7	640.07	-3.1	3.5
Revenues	28	42.58	6.7	15.63	26.0	53.8
Expenditures	28	42.85	6.8	15.85	26.3	56.5

In a metric and classical context, we simply have to apply a standardised PCA on the homogeneous set of the 4 characteristics of the government finance of the EU-28.

In a topological context, the main results of the proposed method are presented in the following tables and graphs, which allow us to visualize proximity measures close to each other and to select the one that best describes and synthesis,

the government finance of the EU-28.

The objective here is to give a topological synthesis of the public finances of the EU countries in 2017.

An HAC algorithm based on the Ward criterion [22], aggregation based on the criterion of the loss of minimal inertia, was used in order to characterize classes of proximity measure relative to their similarities. The

reference measure u^* is projected as a supplementary element. The dendrogram of Figure 2 represents the hierarchical tree of the 15 proximity measures considered. Table 2 describes the final composition of each class of proximity measures, the results of the chosen partition into three homogeneous classes, obtained from the cut of the hierarchical tree of Figure 2.

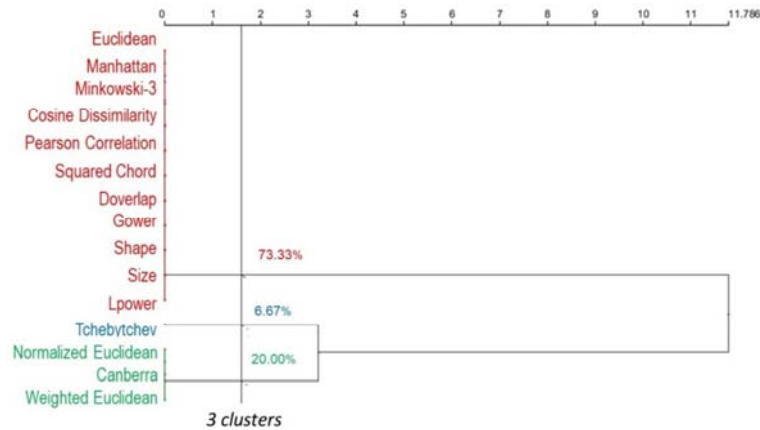


Figure 2. Hierarchical tree of the proximity measures.

Table 2. Clusters composition - Assignment of the reference measure.

Cluster Frequency	Cluster 1 11	Cluster 2, 1	Cluster 3, 3
Proximity measure	Euclidean, Manhattan, Minkowski-3, Cosine, Pearson, Chord, Doverlap, Gower, Shape, Size, Lpower	Tchebychev	Canberra, Normalized Euclidean, Weighted Euclidean
Reference			u^*

Table 3. Similarities and Fisher's Exact Test.

u_i	u_j	$S(u_i, u_j)$	p-value
Cluster 1	Cluster 1	1.000	0.0083**
Cluster 1	Cluster 2	0.750	0.1833
Cluster 1	Cluster 3	0.750	0.1833
Cluster 2	Cluster 2	1.000	0.0083**
Cluster 2	Cluster 3	0.500	1.0000
Cluster 3	Cluster 3	1.000	0.0083**
u^*	Cluster 1	0.750	0.1833
u^*	Cluster 2	0.625	0.5000
u^*	Cluster 3	0.875	0.0333*

Table 4. 2×2 Contingency Table - Similarity - Fisher's Exact Test.

Cluster 2	Cluster 1	Euclidean	Reference	Cluster 1	Euclidean
Tchebychev	$V_{u1}=0$	$V_{u1}=1$	Measure	$V_{u1}=0$	$V_{u1}=1$
$V_{u2}=0$	2	1	$V_{u^*}=0$	3	1
$V_{u2}=1$	1	6	$V_{u^*}=1$	0	6
$S(V_{u2}, V_{u1})=75\%$; p-value=0.183			$S(V_{u^*}, V_{u1})=75\%$; p-value=0.183		

Cluster 3	Cluster 2	Tchebychev	Reference	Cluster 2	Tchebychev
Canberra	$V_{u1}=0$	$V_{u1}=1$	Measure	$V_{u1}=0$	$V_{u1}=1$
$V_{u2}=0$	1	2	$V_{u^*}=0$	2	2
$V_{u2}=1$	2	5	$V_{u^*}=1$	1	5
$S(V_{u3}, V_{u2})=50\%$; p-value=1.000			$S(V_{u^*}, V_{u2})=62.50\%$; p-value=0.500		

Cluster 1	Cluster 3	Canberra	Reference	Cluster 3	Canberra
Euclidean	$V_{u1}=0$	$V_{u1}=1$	Measure	$V_{u1}=0$	$V_{u1}=1$
$V_{u2}=0$	2	1	$V_{u^*}=0$	3	1
$V_{u2}=1$	1	6	$V_{u^*}=1$	0	6
$S(V_{u1}, V_{u3})=75\%$; p-value=0.183			$S(V_{u^*}, V_{u3})=87.50\%$; p-value=0.033*		

Significance level α ; ** $\alpha \leq 1\%$; * $\alpha \in [1\%; 5\%]$

Moreover, in view of the results in Table 2, the reference measure u^* is closer to the third class consisting of Normalized Euclidean, Canberra and Weighted Euclidean measures for which there is a strong topological association between the variables of government finance of EU-28 among the 15 proximity measures considered.

It was shown in [24], by means of a series of experiments, that the choice of proximity measure has an impact on the results of a supervised or unsupervised classification.

In a topological framework, Table 3 summarizes all the results of Table 8 given in the Appendix, the similarities and Fisher's Exact p-values between all the $C_{15}^2=105$ pairs of proximity measures formed with the 15 measures considered and the 15 pairs formed with the unknown reference measure u^* . The values below the diagonal correspond to the similarities $S(V_{ui}, V_{uj})$ and the values above the diagonal are

the Fisher's Exact test p-values.

The similarities in pairs between the 15 proximity measures differ somewhat: some are closer than others, some measures are in perfect topological equivalence $S(V_{ui}, V_{uj})=1$ with a significant Fisher's exact test p-value $< 5\%$; these are therefore identical for the data considered, as is the case with the measures in each cluster of the partition presented in Table 2. The Table 4 illustrates the contingency tables 2×2 between the measures of each cluster: Euclidean, Tchebychev, Canberra and reference measure u^* for the calculation of Fisher's exact test.

Only the topological equivalence between the reference proximity measure and the Canberra proximity measure is significant, $p\text{-value}=0.0034 < \alpha=5\%$, the null hypothesis H_0 of independence is rejected.

Table 5. Pearson correlation matrix (p-value).

Variable	Debt	Deficit	Revenues	Expenditures
Debt	1.000			
Deficit	-0.340 (0.08)	1.000		
Revenues	0.307 (0.11)	0.039 (0.843)	1.000	
Expenditures	0.385 (0.04*)	-0.209 (0.255)	0.969 (0.001**)	1.000

Significance level α ; ** $\alpha \leq 1\%$; * $\alpha \in [1\%; 5\%]$

The adjacency matrix V_{u^*} associated to the adapted proximity measure u^* to the considered data, is build from the correlations matrix Table 5. Figure 5 shows on the main first

TPCA plane, the topological correlation between the Government finance variables.

$$V_{u^*} = \begin{pmatrix} 1 & 0 & 0 & 1 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 1 \\ 1 & 0 & 1 & 1 \end{pmatrix}$$

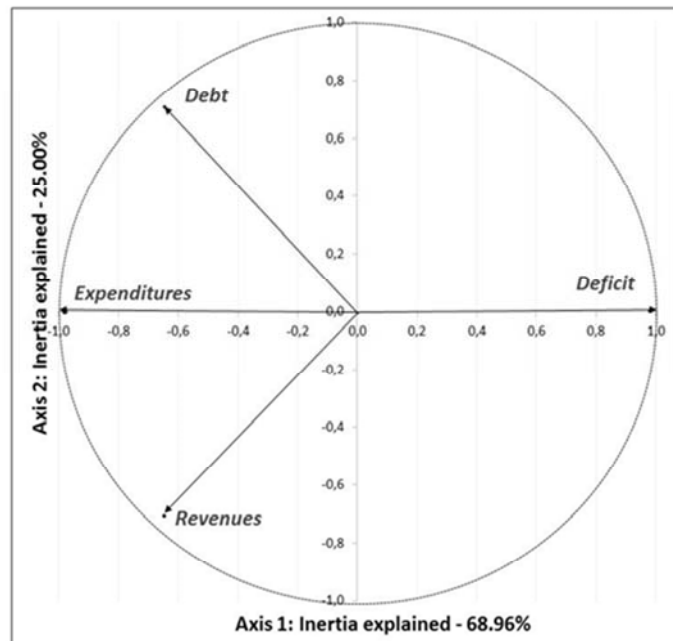


Figure 3. TPCA - Adjacency matrix – The public finance variables on the first principal plane.

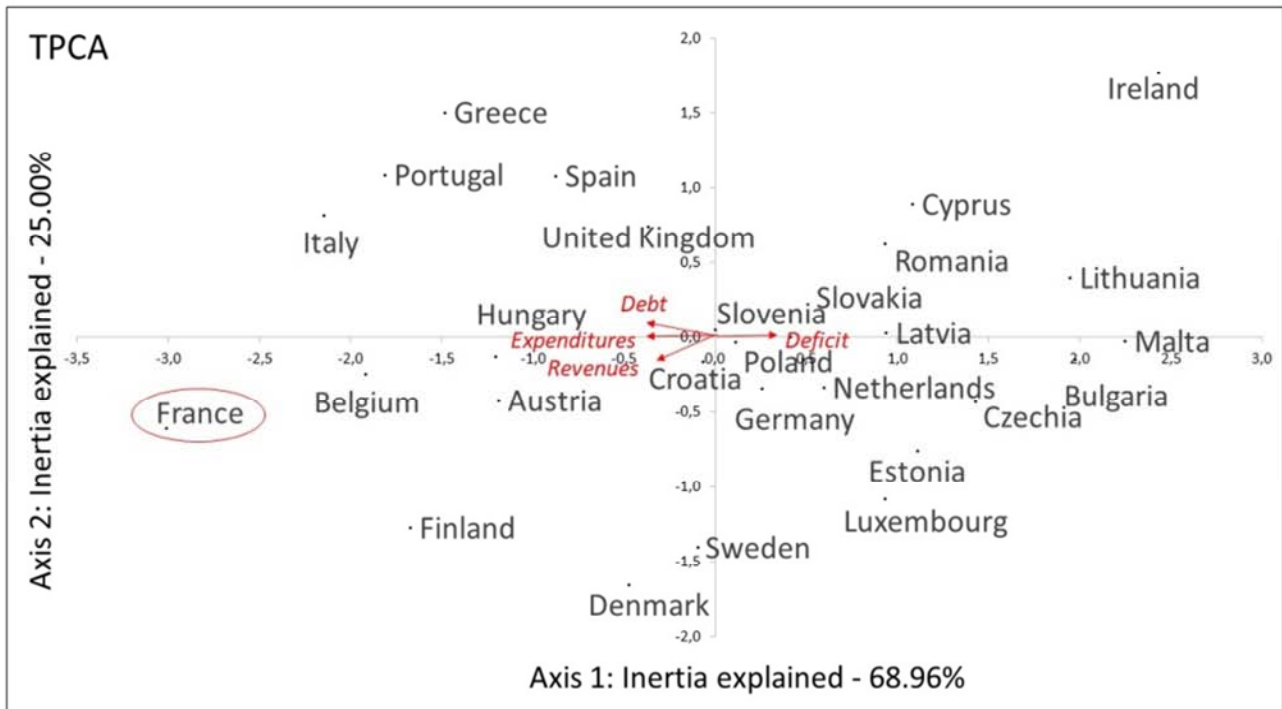


Figure 4. TPCA - The EU-28 countries on the first principal plane.

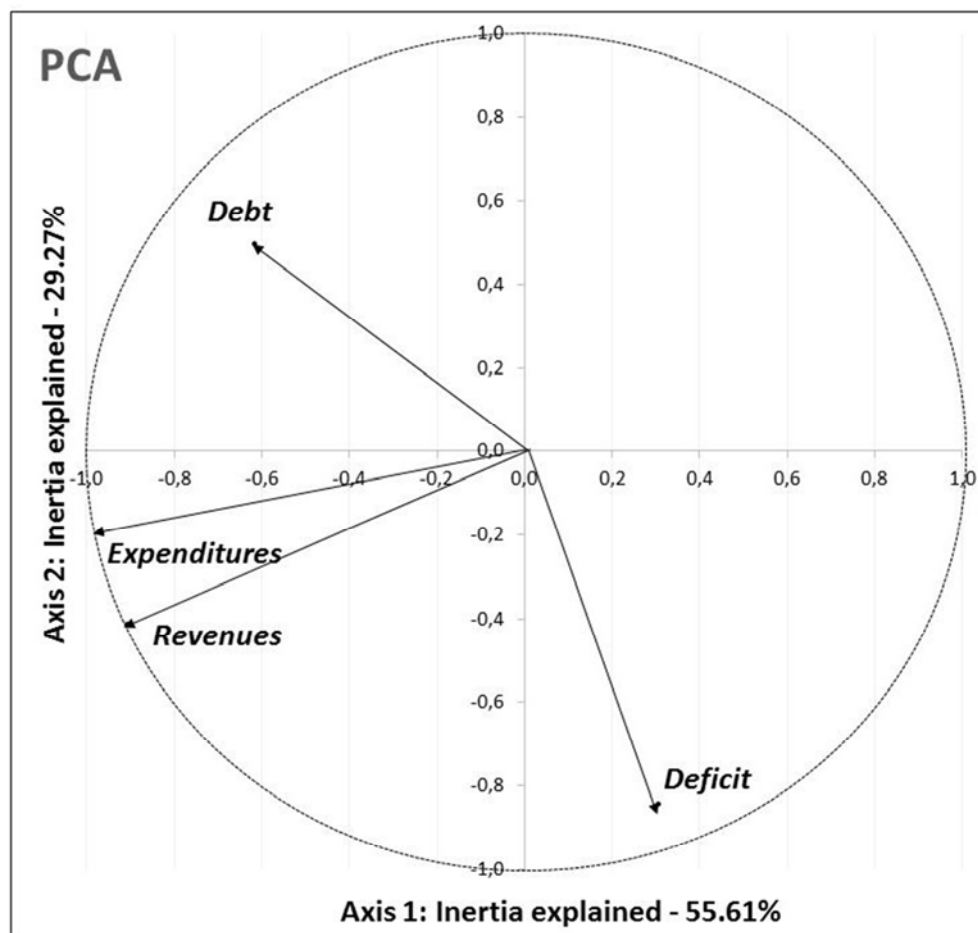


Figure 5. PCA – The public finance variables on the first principal plane.

The adjacency matrix V_u^* associated to the adapted proximity measure u^* to the considered data, is build from the correlation matrix Table 5. Figure 5 shows on the main first TPCA plane, the topological correlation between the Government finance variables.

The corresponding representation for individuals is given in Figure 4. It is thus possible to suggest which are the variables - government finance - responsible for the proximities between the individuals, the 28 EU countries.

The main numerical and graphical results of the proposed TPCA are given in the following Tables and Figures, and are compared to those of the classical PCA.

Figure 5 presents, for comparison on the first factorial plane, the correlations between principal components - Factors and

the original variables. We can see that these graphical representations of the variables are slightly different. Effectively, the percentage of inertia explained on the first principal plane of the Topological PCA is greater than that of classical PCA and the significant correlations variables-factors are also different.

Table 6 shows that the two first factors of TPCA explain 68.96% and 25.00%, respectively, they account for 93.96% of the total variation in the dataset, while the two first factors of classical PCA sum up that 84.88%.

Thus, the first two factors provide an adequate summary of the data, i.e. of government finance of EU-28 countries, we restrict the comparison of the graphical representations to the first factorial plane.

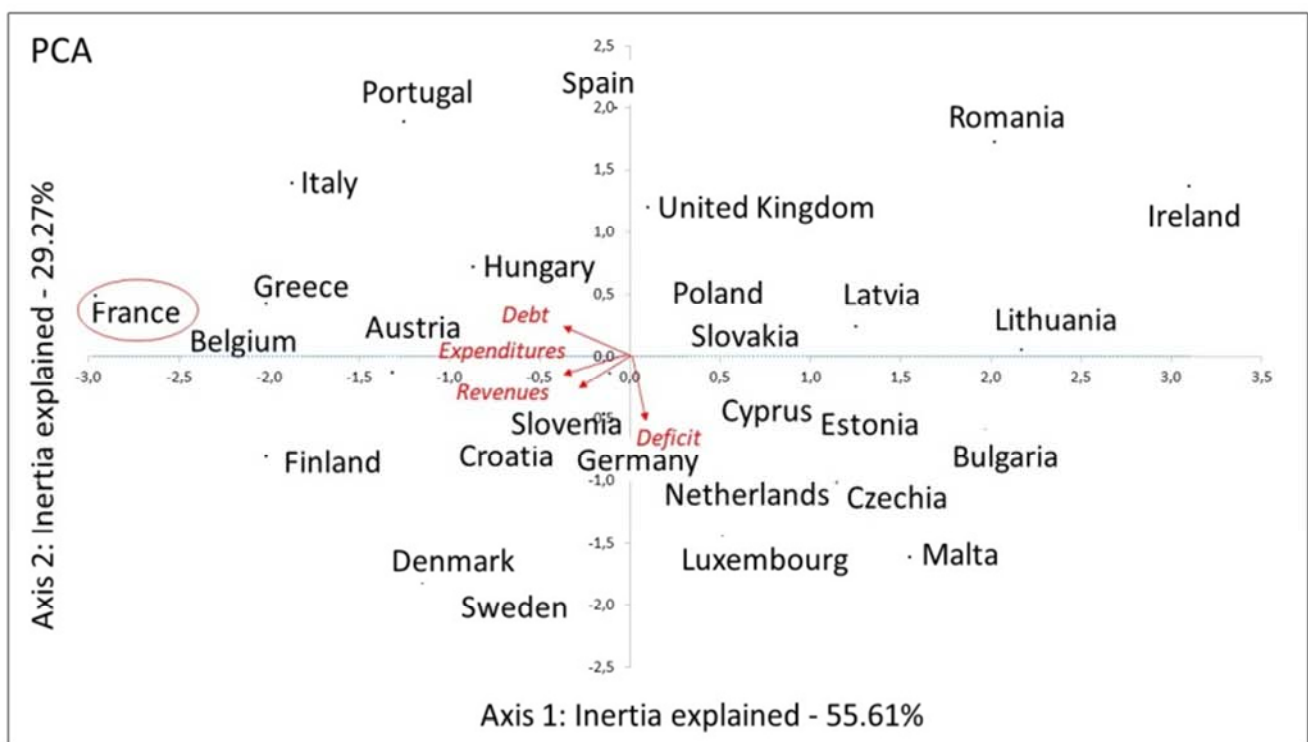


Figure 6. PCA - The EU-28 countries on the first principal plane.

Table 6. TPCA and PCA eigenvalues and correlations Variables & Factors.

TPCA	Eigenvalue	Proportion	Cumulative	Correlation	Factors	
	2.758	68.96%	68.96%	Variables	F1	F2
	1.000	25.00%	93.96%	Debt	0.645	0.707
	0.242	6.04%	100.00%	Deficit	0.982	0.000
	0.000	0.00%	100.00%	Revenues	0.645	-0.707
	4	100.00%	100.00%	Expenditures	0.982	0.000

PCA	Eigenvalue	Proportion	Cumulative	Correlation	Factors	
	2.224	55.61%	55.61%	Variables	F1	F2
	1.171	29.27%	84.88%	Debt	-0.615	0.497
	0.605	15.12%	100.00%	Deficit	0.307	-0.845
	0.000	0.00%	100.00%	Revenues	-0.907	-0.414
	4	100.00%	100.00%	Expenditures	-0.964	-0.196

The correlation tables show that the original variables are strongly correlated with the factors, those that contribute the most to the achievement of this principal component.

While the first PCA factor (55.61%) is strongly correlated with three of the original variables, expenditures, revenues and debt, the first TPCA factor (68.96%) opposes these three variables to the deficit. As for the second PCA (29.27%) and TPCA (25.00%) factors, they oppose the debt to revenues.

The representations of the countries presented in Figures 4 and 6 are of course slightly different, indeed, for example, for France which contributes to the realization of the first TPCA

axis, it is characterized by high Debts, high Expenditures, high Revenues and a low Deficit. France also contributes on the first PCA axis, it's characterized by high Debts, high Expenditures and high Revenues, but the Deficit does not characterize the first factorial axis of the PCA.

We can represent the topological analysis of each of the 15 proximity measures considered, for example see the Euclidean TPCA in Figure 7. One can moreover give Figure 8, the graphical representation associated with a perfect no correlation between variables, from the identity adjacency matrix.

$$V_{u_{Euclidean}} = \begin{pmatrix} 1 & 0 & 0 & 1 \\ 0 & 1 & 1 & 0 \\ 0 & 1 & 1 & 1 \\ 1 & 0 & 1 & 1 \end{pmatrix}$$

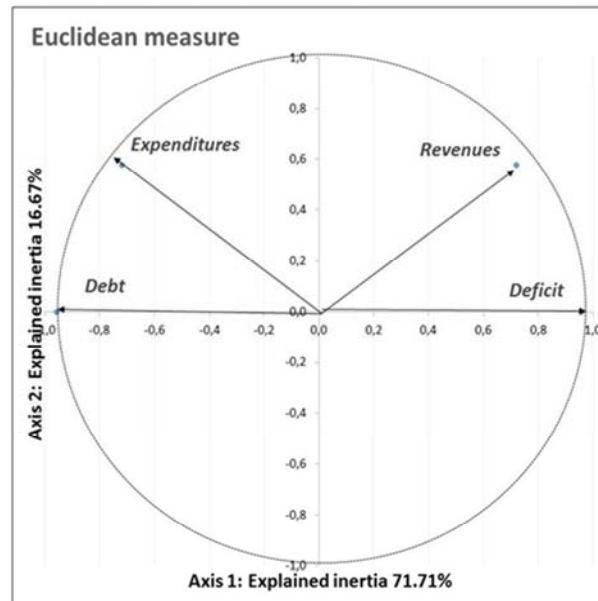


Figure 7. Euclidean TPCA - The public finance variables on the first principal plane.

$$V_{u_{Identity}} = \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

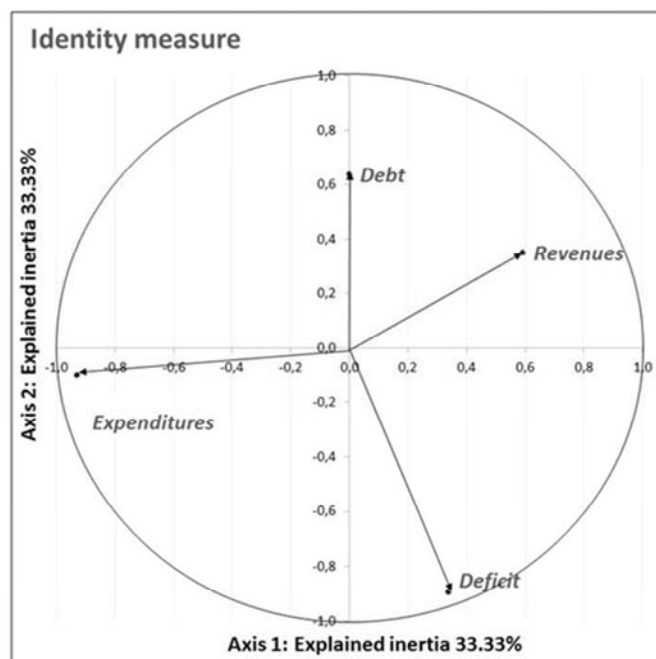


Figure 8. Identity TPCA - The public finance variables on the first principal plane.

4. Conclusion

This research work proposes a new approach that allows to synthesize and describe the correlation structure of a set of quantitative variables in a topological context. Like PCA, the proposed TPCA is a multidimensional topological exploratory method that can be useful for dimension reduction and information redundancy in a data set, it enriches the

conventional quantitative data analysis methods. Future work involves extending this topological approach in three directions, to synthesize the relations existing between a set of a mixture of qualitative and quantitative variables, between two sets of continuous variables in the context of canonical analysis and also between several multidimensional data tables in the context of evolutionary data analysis.

Appendix

Table 7. Similarities & Fisher's Exact Test p-values.

	Euclidean	Manhattan	Minkowski-3	Dissimilarity	Correlation	Squared Chord	Doverlap	Gower
u* measure	0.033	0.033	0.033	0.033	0.033	0.033	0.033	0.033
Euclidean	1	0.008	0.008	0.008	0.008	0.008	0.008	0.008
Manhattan	1	1	0.008	0.008	0.008	0.008	0.008	0.008
Minkowski-3	1	1	1	0.008	0.008	0.008	0.008	0.008
Dissimilarity	1	1	1	1	0.008	0.008	0.008	0.008
Correlation	1	1	1	1	1	0.008	0.008	0.008
Squared Chord	1	1	1	1	1	1	0.008	0.008
Doverlap	1	1	1	1	1	1	1	0.008
Gower	1	1	1	1	1	1	1	1
Shape	1	1	1	1	1	1	1	1
Size	1	1	1	1	1	1	1	1
Lpower	1	1	1	1	1	1	1	1
Tchebytchev	0.75	0.75	0.75	0.75	0.75	0.75	0.75	0.75
N. Euclidean	0.75	0.75	0.75	0.75	0.75	0.75	0.75	0.75
Canberra	0.75	0.75	0.75	0.75	0.75	0.75	0.75	0.75
W. Euclidean	0.75	0.75	0.75	0.75	0.75	0.75	0.75	0.75
u* measure	0.875	0.875	0.875	0.625	0.875	0.875	0.875	0.875

Table 7. Continued.

	Shape	Size	Lpower	Tchebytchev	N. Euclidean	Canberra	W. Euclidean
u* measure	0.033	0.033	0.033	0.5	0.033	0.033	0.033
Euclidean	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Manhattan	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Minkowski-3	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Dissimilarity	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Correlation	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Squared Chord	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Doverlap	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Gower	0.008	0.008	0.008	0.183	0.183	0.183	0.183
Shape	1	0.008	0.008	0.183	0.183	0.183	0.183
Size	1	1	0.008	0.183	0.183	0.183	0.183
Lpower	1	1	1	0.183	0.183	0.183	0.183
Tchebytchev	0.75	0.75	0.75	1	1	1	1
N. Euclidean	0.75	0.75	0.75	0.5	1	0.008	0.008
Canberra	0.75	0.75	0.75	0.5	1	1	0.008
W. Euclidean	0.75	0.75	0.75	0.5	1	1	1
u* measure	0.875	0.875	0.875	0.875	0.875	0.875	0.875

Similarity: S (Tchebytchev; Euclidean)=75%.

Fisher's Exact Test: p-value (Euclidean; Tchebytchev)=0.183 > $\alpha=5\%$: not significant.

Table 8. Some proximity measures for continuous data.

Measure	Formula: Distance - Dissimilarity
Euclidean	$u_{Euclidean}(x, y) = \sqrt{\sum_{j=1}^p (x_j - y_j)^2}$
Manhattan	$u_{Manhattan}(x, y) = \sum_{j=1}^p x_j - y_j $
Minkowski-v	$u_{Minkowski}(x, y) = \left(\sum_{j=1}^p x_j - y_j ^v \right)^{1/v}$
Cosine Dissimilarity	$u_{Cosine}(x, y) = 1 - \frac{\sum_{j=1}^p x_j y_j}{\sqrt{\sum_{j=1}^p x_j^2} \sqrt{\sum_{j=1}^p y_j^2}}$
Pearson Correlation	
Squared Chord	$u_{Chord}(x, y) = \sum_{j=1}^p (\sqrt{x_j} - \sqrt{y_j})^2$
Doverlap measure	$u_{Doverlap}(x, y) = \max \left(\sum_{j=1}^p x_j, \sum_{j=1}^p y_j \right) - \sum_{j=1}^p \min(x_j, y_j)$
Gower	$u_{Gower}(x, y) = \frac{1}{p} \sum_{j=1}^p x_j - y_j $
Shape Distance	$u_{Shape}(x, y) = \sqrt{\sum_{j=1}^p [(x_j - \bar{x}_j) - (y_j - \bar{y}_j)]^2}$
Size Distance	$u_{Size}(x, y) = \left \sum_{j=1}^p (x_j - y_j) \right $
Lpower	$u_{Lpower}(x, y) = \sum_{j=1}^p x_j - y_j ^v$
Tchebychev	$u_{Tchebychev}(x, y) = \max_{1 \leq j \leq p} x_j - y_j $
Normalized Euclidean	$u_{NEuclidean}(x, y) = \sqrt{\sum_{j=1}^p \frac{1}{\sigma_j^2} [(x_j - \bar{x}_j) - (y_j - \bar{y}_j)]^2}$
Canberra	$u_{Canberra}(x, y) = \sum_{j=1}^p \frac{ x_j - y_j }{ x_j + y_j }$
Weighted Euclidean	$u_{WEuclidean}(x, y) = \sqrt{\sum_{j=1}^p \alpha_j (x_j - y_j)^2}$

Where, p is the dimension of space, $x=(x_j)_{j=1, \dots, p}$ and $y=(y_j)_{j=1, \dots, p}$ two points in R^p , \bar{x}_j the mean, σ_j the Standard deviation, $\alpha_j=1/\sigma_j^2$ and $v > 0$.

References

- [1] R. Abdesselam, "A Topological Multiple Correspondence Analysis." Journal of Mathematics and Statistical Science, Science Signpost Publishing Inc., USA, Vol. 5, Issue 8, pp. 175-192, 2019.
- [2] R. Abdesselam, "Selection of proximity measures for a Topological Correspondence Analysis." In a Book Series, 5th Stochastic Modeling Techniques and Data Analysis, International Conference, Chania, Greece, pp. 11-24, 2018.
- [3] R. Abdesselam, "A Topological Discriminant Analysis." In book Chapter, Vol. 3, Data Analysis and Applications 2: Utilization of Results in Europe and Other Topics, ISTE Science Publishing, Wiley, pp. 167-178, 2018.
- [4] V. Batagelj and M. Bren, "Comparing resemblance measures." In Journal of classification, 12, pp. 73-90, 1995.
- [5] F. Cailliez and J-P Pagès "Introduction à l'Analyse des données.", S. M. A. S. H., Paris, 1976.
- [6] J. Cohen, "A coefficient of agreement for nominal scales." Educational and Psychological Measurement, Vol. 20, pp. 27-46, 1960.
- [7] J. Demsar, "Statistical comparisons of classifiers over multiple data sets." The journal of Machine Learning Research, Vol. 7, pp. 1-30, 2006.
- [8] Eurostat, Data source: Government finance statistics - Statistics explained, https://ec.europa.eu/eurostat/statistics-explained/index.php?title=Government_finance_statistics, pp. 1-15, 2018.
- [9] R-A. Fisher, "The Interpretation of chi2 from Contingency Tables, and the Calculation of P." Journal of the Royal Statistical Society, Published by Wiley, 85, 1, pp. 87-94, 1922.
- [10] H. Hotelling, "Analysis of a Complex of Statistical Variables into Principal Components". Journal of Educational Psychology, Vol. 24, 6, pp. 417-441, 1933.
- [11] J. H. Kim and S. Lee, "Tail bound for the minimal spanning tree of a complete graph." In Statistics & Probability Letters, Vol. 4, 64, pp. 425-430, 2003.
- [12] L. Lebart, "Stratégies du traitement des données d'enquêtes." La Revue de MODULAD, 3, pp. 21-29, 1989.
- [13] J. Lesot, M. Rifqi and H. Benhadda, "Similarity measures for binary and numerical data: a survey." In IJKE SDP, Vol. 1, 1, pp. 63-84, 2009.

- [14] N. Mantel, "A technique of disease clustering and a generalized regression approach." In *Cancer Research*, Vol. 27, pp. 209-220, 1967.
- [15] J-C. Park, H. Shin, and B-K. Choi, "Elliptic Gabriel graph for finding neighbors in a point set and its application to normal vector estimation." In *Computer-Aided Design Elsevier*, Vol. 38, 6, pp. 619-626, 2006.
- [16] K. Pearson, "On lines and Planes of Closest Fit to Systems of Points in Space." In *Philosophical Magazine*, vol. 2, 11, pp. 559-572, 1901.
- [17] J. Rifqi, M., Detyniecki, M. and Bouchon-Meunier, B. "Discrimination power of measures of resemblance." IFSA'03 Citeseer, 2003.
- [18] G. Saporta, "Probabilités, analyse des données et Statistique." Editions TECHNIP, 2011.
- [19] J-W. Schneider and P. Borlund, "Matrix comparison, Part 1: Motivation and important issues for measuring the resemblance between proximity measures or ordination results." In *Journal of the American Society for Information Science and Technology*, Vol. 58, 11, pp. 1586-1595, 2007.
- [20] J-W. Schneider and P. Borlund, "Matrix comparison, Part 2: Measuring the resemblance between proximity measures or ordination results by use of the Mantel and Procrustes statistics." In *Journal of the American Society for Information Science and Technology*, Vol. 11, 58, pp. 1596-1609, 2007.
- [21] G-T. Toussaint, "The relative neighbourhood graph of a finite planar set." In *Pattern recognition*, Vol. 12, 4, pp. 261-268, 1980.
- [22] J-R. Ward, "Hierarchical grouping to optimize an objective function." In *Journal of the American statistical association JSTOR*, Vol. 58, 301, pp. 236-244, 1963.
- [23] M-J. Warrens, "Bounds of resemblance measures for binary (presence/absence) variables." In *Journal of Classification*, Springer, Vol. 25, 2, pp. 195-208, 2008.
- [24] D. Zighed, R. Abdesselam and A. Hadgu, "Topological comparisons of proximity measures." In the 16th PAKDD 2012 Conference. In P.-N. Tan et al., Eds. Part I, LNAI 7301, Springer-Verlag, Berlin Heidelberg, pp. 379-391, 2012.